## Clustering above Exponential Families with Tempered Exponential Measures

Ehsan Amid, Richard Nock and Manfred K. Warmuth

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<u>Why this work</u>: *k*-means = popular clustering, partition of space  $\mathfrak{X} \subseteq \mathbb{R}^d$  by finding set of centers  $\mathfrak{C} \doteq {\mathbf{c}_j}_{j \in [k]}$  minimizing loss to *m*-sample,  $\mathbb{E}_{i \sim [m]} [\min_{j \in [k]} D(\theta_i \| \mathbf{c}_j)]$ 

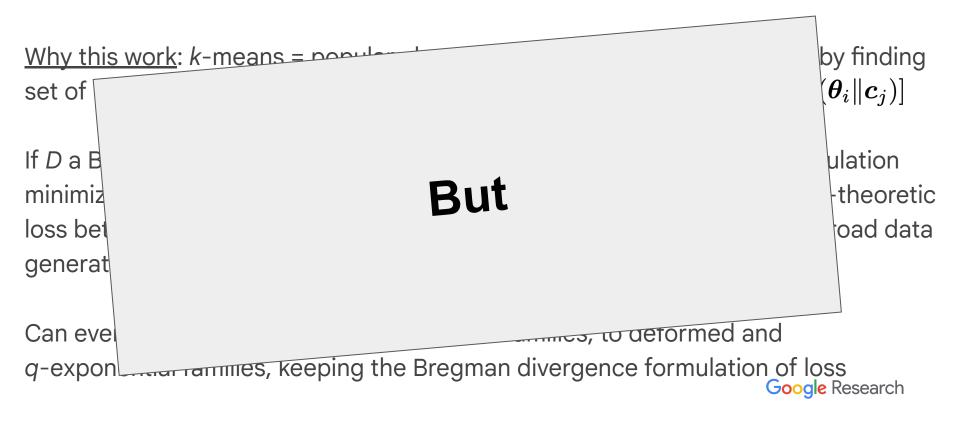
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If *D* a Bregman divergence (Mahalanobis, Itakura-Saito, KL, etc.) then population minimizers trivial to compute **and** we equivalently minimise an information-theoretic loss between *distributions* in exponential families  $\Rightarrow$  embeds *k*-means in broad data generating processes

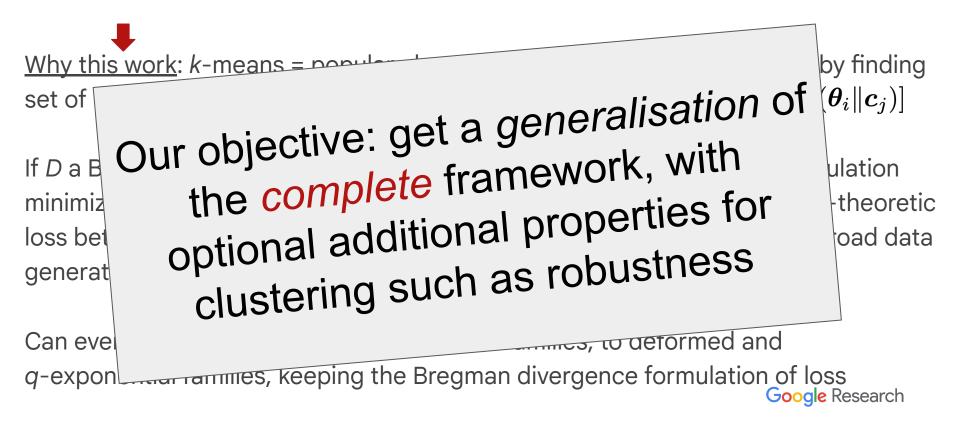
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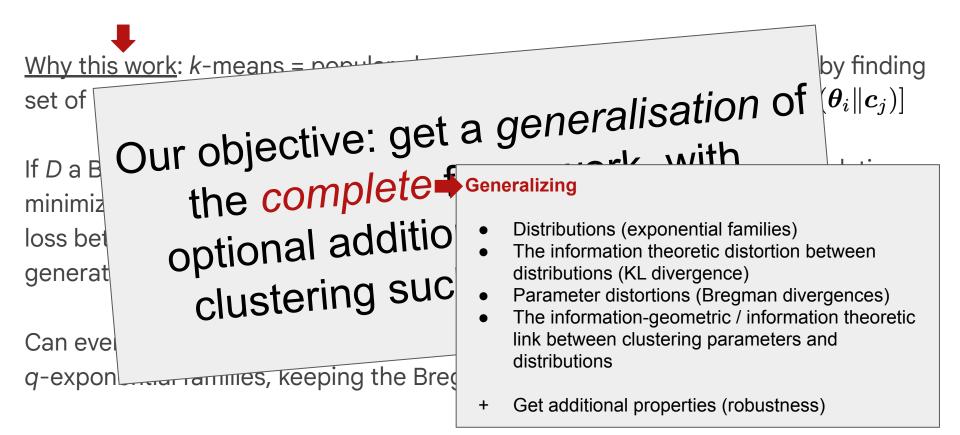
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Can even be generalized above exponential families, to deformed and *q*-exponential families, keeping the Bregman divergence formulation of loss Google Research



Why this	Work: k-means = populated Universal modeling with Bregman	by finding
set of	Universal modeling with bieg	$(oldsymbol{ heta}_i \  oldsymbol{c}_j)]$
lf D a B		ulation
minimiz		theoretic
loss bet		oad data
generat	drawbacks, such do low robustness to outliers (a Bregman robustness)	
Can eve	robustness to outlione ( divergence lacks robustness)	
	rannes, keeping the Bregman divergence formulation of lo	DSS e Research





#### Axiomatic characterization

Set of probability measures satisfying a constraint on their expectation

$$ilde{\mathcal{P}}_{t|\boldsymbol{\hbar}} \doteq egin{cases} & \left. \mathbb{E}_{ ilde{P}}[\boldsymbol{\phi}] \doteq \int \boldsymbol{\phi}(\boldsymbol{x}) \, ilde{p}(\boldsymbol{x}) \, \mathrm{d}\xi = \boldsymbol{\hbar}, \ & \int ilde{p}(\boldsymbol{x}) \, \mathrm{d}\xi = 1, \ & ilde{p}(\boldsymbol{x}) \geq 0, orall \boldsymbol{x} \in \mathfrak{X}. \end{aligned} 
ight.$$

+constraint to maximize entropy  $H(P) \doteq -\int p \log p d\xi$  $\Rightarrow$  get an exponential family

$$p_{oldsymbol{ heta}}(oldsymbol{x}) \propto \exp(oldsymbol{ heta}^ op \phi(oldsymbol{x}) - G(oldsymbol{ heta}))$$
 $\stackrel{\uparrow}{oldsymbol{ heta}} = 
abla G^{-1}(oldsymbol{\hbar})$ 
 $\operatorname{ratural parameter}$ 
 $\operatorname{cumulant}$ 

Google Research

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Set of unnormalized measures satisfying

a constraint on their expectation

$$\tilde{\mathcal{P}}_{t|\boldsymbol{\hbar}} \doteq \left\{ \tilde{p} \middle| \begin{array}{c} \mathbb{E}_{\tilde{P}}[\boldsymbol{\phi}] \doteq \int \boldsymbol{\phi}(\boldsymbol{x}) \, \tilde{p}(\boldsymbol{x}) \, \mathrm{d}\boldsymbol{\xi} = \boldsymbol{\hbar}, \\ \int \tilde{p}(\boldsymbol{x})^{1/t^*} \, \mathrm{d}\boldsymbol{\xi} = 1, \\ \tilde{p}(\boldsymbol{x}) \ge 0, \forall \boldsymbol{x} \in \boldsymbol{\mathfrak{X}}. \end{array} \right\}$$

$$t \in [0,1], t^* \doteq 1/(2-t)$$

#### Axiomatic characterization

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$$p_{\theta}(\boldsymbol{x}) \propto \exp(\boldsymbol{\theta}^{\top} \boldsymbol{\phi}(\boldsymbol{x}) - G(\boldsymbol{\theta}))$$
  
 $\stackrel{\uparrow}{\boldsymbol{\theta}} = \nabla G^{-1}(\boldsymbol{\hbar})$  cumulant  
natural parameter

Set of **unnormalized** measures satisfying a constraint on their expectation

$$\tilde{\mathcal{P}}_{t|\hbar} \doteq \left\{ \tilde{p} \middle| \begin{array}{c} \mathbb{E}_{\tilde{P}}[\boldsymbol{\phi}] \doteq \int \boldsymbol{\phi}(\boldsymbol{x}) \, \tilde{p}(\boldsymbol{x}) \, \mathrm{d}\boldsymbol{\xi} = \boldsymbol{\hbar}, \\ \int \tilde{p}(\boldsymbol{x})^{1/t^*} \, \mathrm{d}\boldsymbol{\xi} = 1, \\ \tilde{p}(\boldsymbol{x}) \ge 0, \forall \boldsymbol{x} \in \boldsymbol{\mathfrak{X}}. \end{array} \right\}$$

+maximize a generalized Tsallis entropy  $H_t(\tilde{P}) \doteq -\int (\tilde{p}\log_t \tilde{p} - \log_{t-1} \tilde{p})d\xi$ tempered log  $\longrightarrow \log_t(z) \doteq \frac{1}{1-t} (z^{1-t} - 1)$ 

Concave,  $\lim_{t \to 1} \log_t = \log_t$ 

#### Axiomatic characterization

Set of probability measures satisfying a constraint on their expectation

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$$p_{oldsymbol{ heta}}(oldsymbol{x}) \propto \exp(oldsymbol{ heta}^{ op} oldsymbol{\phi}(oldsymbol{x}) - G(oldsymbol{ heta}))$$
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abla G^{-1}(oldsymbol{\hbar})$ 
 $\operatorname{cumulan}$ 
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Set of unnormalized measures satisfying a constraint on their expectation

$$\tilde{\mathcal{P}}_{t|\boldsymbol{\hbar}} \doteq \left\{ \tilde{p} \middle| \begin{array}{c} \mathbb{E}_{\tilde{P}}[\boldsymbol{\phi}] \doteq \int \boldsymbol{\phi}(\boldsymbol{x}) \, \tilde{p}(\boldsymbol{x}) \, \mathrm{d}\boldsymbol{\xi} = \boldsymbol{\hbar}, \\ \int \tilde{p}(\boldsymbol{x})^{1/t^*} \, \mathrm{d}\boldsymbol{\xi} = 1, \\ \tilde{p}(\boldsymbol{x}) \ge 0, \forall \boldsymbol{x} \in \boldsymbol{\mathfrak{X}}. \end{array} \right\}$$

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 $\Rightarrow \text{get a tempered exponential measure}$   $\tilde{p}_{t|\theta}(\boldsymbol{x}) \propto \frac{\exp_t(\boldsymbol{\theta}^{\top}\boldsymbol{\phi}(\boldsymbol{x}))}{\exp_t(G_t(\boldsymbol{\theta}))}$   $\stackrel{\uparrow}{=} \nabla G_t^{-1}(\boldsymbol{\hbar}) \qquad \text{cumulant}$ 

$$\widetilde{p}_{t|\boldsymbol{\theta}}(\boldsymbol{x}) \propto \exp_t(\boldsymbol{\theta}^{\top} \boldsymbol{\phi}(\boldsymbol{x}) \ominus_t G_t(\boldsymbol{\theta}))$$
 $z \ominus_t x \doteq \frac{z - x}{1 + (1 - t)x}$ 

#### **Tempered Exponential Measures**

Set of unnormalized measures satisfying a constraint on their expectation

$$\tilde{\mathcal{P}}_{t|\boldsymbol{\hbar}} \doteq \left\{ \tilde{p} \middle| \begin{array}{c} \mathbb{E}_{\tilde{P}}[\boldsymbol{\phi}] \doteq \int \boldsymbol{\phi}(\boldsymbol{x}) \, \tilde{p}(\boldsymbol{x}) \, \mathrm{d}\boldsymbol{\xi} = \boldsymbol{\hbar}, \\ \int \tilde{p}(\boldsymbol{x})^{1/t^*} \, \mathrm{d}\boldsymbol{\xi} = 1, \\ \tilde{p}(\boldsymbol{x}) \ge 0, \forall \boldsymbol{x} \in \boldsymbol{\mathfrak{X}}. \end{array} \right\}$$

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⇒ get a tempered exponential measure

$$\widetilde{p}_{t|\boldsymbol{ heta}}(\boldsymbol{x}) \propto rac{\exp_t(\boldsymbol{ heta}^{ op}\boldsymbol{\phi}(\boldsymbol{x}))}{\exp_t(G_t(\boldsymbol{ heta}))_{ ext{Google Research}}}$$
  
 $\boldsymbol{ heta} = \nabla G_t^{-1}(\boldsymbol{\hbar})$ 
cumulant

$$\tilde{p}_{t|\boldsymbol{\theta}}(\boldsymbol{x}) \propto \exp_t(\boldsymbol{\theta}^{\top} \boldsymbol{\phi}(\boldsymbol{x}) \ominus_t G_t(\boldsymbol{\theta}))$$
  
 $z \ominus_t x \doteq \frac{z-x}{1+(1-t)x}$ 

Cumulant in closed form:

$$G_t(\boldsymbol{\theta}) = (\log_t)^* \int (\exp_t)^* (\boldsymbol{\theta}^\top \boldsymbol{\phi}(\boldsymbol{x})) d\xi$$
$$(\log_t)^* (z) \doteq t^* \log_{t^*} \left(\frac{z}{t^*}\right)$$
$$(\exp_t)^* (z) \doteq t^* \exp_{t^*} \left(\frac{z}{t^*}\right)$$
$$\exp_t(z) \doteq [1 + (1 - t)z]_+^{1/(1 - t)} \text{ tempered exp}$$
$$[.]_+ \doteq \max\{0, .\} \qquad \lim_{t \to 1} \exp_t = \exp_t(z)$$

#### **Tempered Exponential Measures**

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Total mass in closed form:

$$\int \tilde{p}_{t|\boldsymbol{\theta}}(\boldsymbol{x}) \mathrm{d}\boldsymbol{\xi} = 1 + (1-t)(G_t(\boldsymbol{\theta}) - \boldsymbol{\theta}^{\top}\boldsymbol{\hbar})$$

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 $\tilde{p}_{t|\boldsymbol{\theta}}^{1/t^*} \text{= co-density}$ 

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#### Information Geometric Distortions (gen. Bregman divs)

Information-theoretic distortion between two TEMs, generalizing (reverse) KL divergence:

$$F_{t}(\tilde{P}_{t|\hat{\theta}} \| \tilde{P}_{t|\theta}) \doteq \int f\left(\frac{\mathrm{d}\tilde{p}_{t|\hat{\theta}}}{\mathrm{d}\xi} \oslash_{t} \frac{\mathrm{d}\tilde{p}_{t|\theta}}{\mathrm{d}\xi}\right) \cdot \mathrm{d}\tilde{p}_{t|\theta} \qquad \qquad f \doteq -\log_{t} \\ x \oslash_{t} y \doteq [x^{1-t} - y^{1-t} + 1]_{+}^{\frac{1}{1-t}}$$

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Theorem: for any 2 members of the same TEM family,  $F_t(\tilde{P}_{t|\hat{\theta}} \| \tilde{P}_{t|\theta}) = B_{G_t}(\hat{\theta} \| \theta)$ , with

$$B_{G_t}(\hat{\boldsymbol{\theta}} \| \boldsymbol{\theta}) \doteq \frac{G_t(\hat{\boldsymbol{\theta}}) - G_t(\boldsymbol{\theta}) - (\hat{\boldsymbol{\theta}} - \boldsymbol{\theta})^\top \nabla G_t(\boldsymbol{\theta})}{1 + (1 - t)G_t(\hat{\boldsymbol{\theta}})} \ \ \text{Bregman divergence}$$

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#### **Clustering: population minimizers**

Given training sample  $\{\boldsymbol{\theta}_i\}_{i=1}^m$ , we seek its left and right population minimizers, i.e. having  $L_1(\boldsymbol{\theta}) \doteq \mathbb{E}_i[B_{G_t}(\boldsymbol{\theta} \| \boldsymbol{\theta}_i)]$ ;  $L_r(\boldsymbol{\theta}) \doteq \mathbb{E}_i[B_{G_t}(\boldsymbol{\theta}_i \| \boldsymbol{\theta})]$ , we want to compute  $\boldsymbol{\theta}_l \doteq \arg\min_{\boldsymbol{\theta}} L_l(\boldsymbol{\theta})$ ;  $\boldsymbol{\theta}_r \doteq \arg\min_{\boldsymbol{\theta}} L_r(\boldsymbol{\theta})$ 

left population minimizer

right population minimizer

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$$\boldsymbol{\theta}_{\mathrm{l}} \doteq \arg\min_{\boldsymbol{\theta}} L_{\mathrm{l}}(\boldsymbol{\theta}) \quad ; \quad \boldsymbol{\theta}_{\mathrm{r}} \doteq \arg\min_{\boldsymbol{\theta}} L_{\mathrm{r}}(\boldsymbol{\theta})$$

left population minimizer

right population minimizer

Theorem: we have

$$\boldsymbol{\theta}_{\mathrm{l}} = \nabla G_t^{-1}(\alpha_* \cdot \mathbb{E}_i \nabla G_t(\boldsymbol{\theta}_i))$$

$$\boldsymbol{\theta}_{\mathrm{r}} = \mathbb{E}_{i} \left[ \frac{1}{\exp_{t}^{1-t}(G_{t}(\boldsymbol{\theta}_{i}))} \cdot \boldsymbol{\theta}_{i} 
ight]$$

 $\alpha_* > 0$ 

Precise interval to search (*Cf* paper) Closed forms available in particular cases

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#### Clustering: population minimizers... and robustness

Given training sample  $\{\boldsymbol{\theta}_i\}_{i=1}^m$ , we seek its left and right population minimizers, i.e. having  $L_1(\boldsymbol{\theta}) \doteq \mathbb{E}_i[B_{G_t}(\boldsymbol{\theta} \| \boldsymbol{\theta}_i)]$ ;  $L_r(\boldsymbol{\theta}) \doteq \mathbb{E}_i[B_{G_t}(\boldsymbol{\theta}_i \| \boldsymbol{\theta})]$ , we want to compute  $\boldsymbol{\theta}_l \doteq \arg\min_{\boldsymbol{\theta}} L_1(\boldsymbol{\theta})$ ;  $\boldsymbol{\theta}_r \doteq \arg\min_{\boldsymbol{\theta}} L_r(\boldsymbol{\theta})$ 

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Robustness: add outlier  $\boldsymbol{\theta}_*$  with weight  $\boldsymbol{\epsilon}$ . The center moves as  $\boldsymbol{\theta}_{l/r}^{new} - \boldsymbol{\theta}_{l/r}^{old} = \boldsymbol{\epsilon} \cdot \boldsymbol{z}(\boldsymbol{\theta}_*)$ If the influence function,  $\boldsymbol{z}(.)$ , has bounded norm, then the center is robust

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#### Clustering: population minimizers... and robustness

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m l}\doteq rg\min_{oldsymbol{ heta}} L_{
m l}(oldsymbol{ heta}) ~~;~~oldsymbol{ heta}_{
m r}\doteq rg\min_{oldsymbol{ heta}} L_{
m r}(oldsymbol{ heta})$ right population minimizer left population minimizer Theorem: we have  $\boldsymbol{\theta}_{\mathrm{l}} = \nabla G_{t}^{-1}(\alpha_{*} \cdot \mathbb{E}_{i} \nabla G_{t}(\boldsymbol{\theta}_{i})) \qquad \boldsymbol{\theta}_{\mathrm{r}} = \mathbb{E}_{i} \left| \frac{1}{\exp^{1-t}(G_{t}(\boldsymbol{\theta}_{i}))} \cdot \boldsymbol{\theta}_{i} \right|$ Robustness: add outlier  $\boldsymbol{\theta}_*$  with weight  $\boldsymbol{\epsilon}$ . The center moves as  $\boldsymbol{\theta}_{l/r}^{new} - \boldsymbol{\theta}_{l/r}^{old} = \boldsymbol{\epsilon} \cdot \boldsymbol{z}(\boldsymbol{\theta}_*)$ If the influence function, z(.), has bounded norm, then the center is robust

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Theorem: left robust *iff* robust for t = 1; right robust *iff* 

TEM	Support	$\lambda$	θ	ħ	$G_t^\star(\hbar)$
1D $t$ -exponential	$\left[0,rac{3-2t}{(1-t)\lambda} ight]$	$\lambda$	$\frac{-\lambda}{3-2t}$	$t^*\left(rac{3-2t}{\lambda} ight)^{2-t^*}$	$-t^* \cdot \left( \log_{\frac{1}{2-t^*}} \left( \frac{\hbar}{t^*} \right) - 1 \right)$
1D t-Gaussian ( $\mu = 0$ )	$\left[-rac{1}{\sqrt{1-t}},rac{1}{\sqrt{1-t}} ight]$	$\sigma^2$	$-rac{t^{*}}{2\sigma^{2}}$	$(c_{t^*}\sqrt{2})^{1-t^*}\sigma^{3-t^*}$	$-\frac{t^*}{2} \cdot \left(\log_{t^{**}}(2c_{t^*}^2\hbar) - 1\right)$

TEM	Support $\lambda$ $\ell$		θ	ħ	$G_t^\star(\hbar)$	
1D $t$ -exponential	$\left[0,rac{3-2t}{(1-t)\lambda} ight]$	$\lambda$	$rac{-\lambda}{3-2t}$	$t^* \left(rac{3-2t}{\lambda} ight)^{2-t^*}$	$-t^* \cdot \left( \log_{\frac{1}{2-t^*}} \left( \frac{\hbar}{t^*} \right) - 1 \right)$	
1D <i>t</i> -Gaussian $(\mu = 0)$	$\left[-\frac{1}{\sqrt{1-t}},\frac{1}{\sqrt{1-t}}\right]$	$\sigma^2$	$-rac{t^*}{2\sigma^2}$	$(c_{t^*}\sqrt{2})^{1-t^*}\sigma^{3-t^*}$	$-\frac{t^*}{2} \cdot \left( \log_{t^{**}} (2c_{t^*}^2 \hbar) - 1 \right)$	
$\mathbf{TEM}$	$G_t(oldsymbol{ heta})$			$\ B_{G_t}(\hat{oldsymbol{ heta}}\ oldsymbol{ heta})$		
1D t-exponential	$-\log_{2-t}\left((-\theta)^{\frac{1}{2-t}}\right)$		)	$t^* \cdot \left( \left( \frac{\hat{\theta}}{\theta} \right)^{2-t^*} - (2-t^*) \cdot \log_{t^*} \left( \frac{\hat{\theta}}{\theta} \right) - 1 \right)$		
1D <i>t</i> -Gaussian ( $\mu = 0$ )	$\left(\log_t\right)^* \left(\frac{c_{t^*}}{\sqrt{-\theta}}\right)$		$\frac{t^{*}}{2}$	$\frac{t^*}{2} \cdot \left( \left( \sqrt{\frac{\hat{\theta}}{\theta}} \right)^{3-t^*} - (3-t^*) \cdot \log_{t^*} \sqrt{\frac{\hat{\theta}}{\theta}} - 1 \right)$		

Two distinct generalisations of Itakura-Saito divergence !

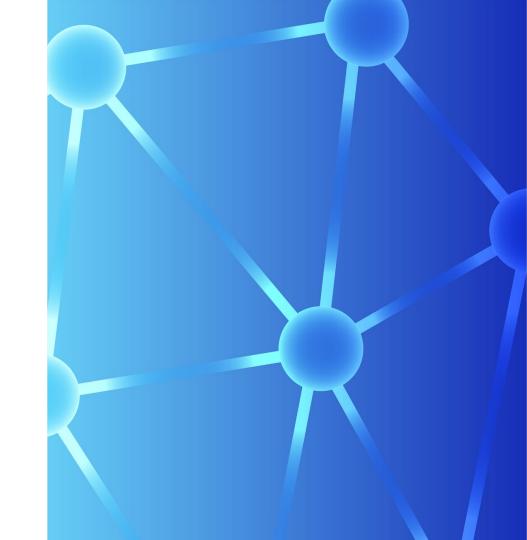
TEM	Support	$\lambda$ (	θ ħ	,	$G_t^{\star}(\hbar)$
1D $t$ -exponential	$\left[0, rac{3-2t}{(1-t)\lambda} ight]$	$\lambda = \frac{-}{3}$	$\frac{-\lambda}{-2t}$ $t^*\left(\frac{3-2}{\lambda}\right)$	$\left(\frac{t}{t}\right)^{2-t^*}$	$-t^* \cdot \left( \log_{\frac{1}{2-t^*}} \left( \frac{\hbar}{t^*} \right) - 1 \right)$
1D <i>t</i> -Gaussian ( $\mu = 0$ )	$\left[-rac{1}{\sqrt{1-t}},rac{1}{\sqrt{1-t}} ight]$	$\sigma^2 - \frac{1}{2}$	$\frac{t^*}{2\sigma^2}  (c_{t^*}\sqrt{2})^{1-1}$	$-t^*\sigma^{3-t^*}$	$-\frac{t^*}{2} \cdot \left(\log_{t^{**}}(2c_{t^*}^2\hbar) - 1\right)$
			-		
TEM	$G_t(\boldsymbol{ heta})$		$B_{G_t}(\hat{oldsymbol{ heta}} \  oldsymbol{ heta})$		
1D $t$ -exponential	$-\log_{2-t}\left((-\theta)^{\frac{1}{2-t}}\right)$		$t^* \cdot \left( \left( \frac{\hat{\theta}}{\theta} \right)^{2-t^*} - (2-t^*) \cdot \log_{t^*} \left( \frac{\hat{\theta}}{\theta} \right) - 1 \right)$		
1D t-Gaussian ( $\mu = 0$ )	$\left(\log_t\right)^* \left(\frac{c_{t^*}}{\sqrt{-\theta}}\right)$		$\frac{t^*}{2} \cdot \left( \left( \sqrt{\frac{\hat{\theta}}{\theta}} \right)^{3-t^*} - (3-t^*) \cdot \log_{t^*} \sqrt{\frac{\hat{\theta}}{\theta}} - 1 \right)$		
TEM	$\boldsymbol{\theta}_1$	$\theta_{\rm r}$			
1D t-exponential	$-\mathbb{E}_{i}\left[\frac{1}{(-\theta_{i})^{1}}\right]$		$\left[\frac{1}{(-\theta_i)^{2-t^*}}\right]$		$\frac{1}{-\mathbb{E}_i\left[(-\theta_i)^{2-t^*}\right]}$
1D <i>t</i> -Gaussian $(\mu = 0)$		$\left[\frac{-t^{*}}{2}\right]/\mathbb{E}_{i}$		$-\frac{1}{(c_t*)}$	$rac{1}{(t^*)^{1-t^*}} \cdot \mathbb{E}_i\left[(- heta_i)^{rac{3-t^*}{2}} ight]$

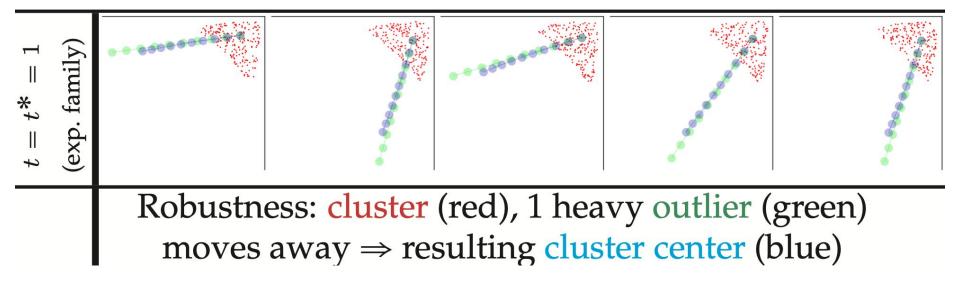
TEM	Support $\lambda$	θ ħ	į	$G^{\star}_t(\hbar)$	
1D $t$ -exponential	$\left[0, \frac{3-2t}{(1-t)\lambda}\right] \qquad \lambda \qquad \frac{3}{3}$	$\frac{-\lambda}{3-2t}$ $t^*\left(\frac{3-2}{\lambda}\right)$	$\left(\frac{t}{2}\right)^{2-t^*}$	$-t^* \cdot \left( \log_{\frac{1}{2-t^*}} \left( \frac{\hbar}{t^*} \right) - 1 \right)$	
1D <i>t</i> -Gaussian $(\mu = 0)$	$-\frac{1}{\sqrt{1-t}},\frac{1}{\sqrt{1-t}}\right]  \sigma^2  -\frac{t^*}{2\sigma^2}  (c_t \cdot \sqrt{2})$		$-t^*\sigma^{3-t^*}$	$-\frac{t^{*}}{2} \cdot \left( \log_{t^{**}} (2c_{t^{*}}^{2}\hbar) - 1 \right)$	
TEM	$G_t(\boldsymbol{ heta})$		$B_{G_t}(\hat{\boldsymbol{ heta}} \  \boldsymbol{ heta})$		
1D $t$ -exponential	$-\log_{2-t}\left((-\theta)^{\frac{1}{2-t}}\right)$	$t^* \cdot \left( \left( \begin{array}{c} 1 \\ Right population minimizer: not the \end{array} \right) \right)$			
1D <i>t</i> -Gaussian $(\mu = 0)$	$\left(\log_t\right)^* \left(\frac{c_{t^*}}{\sqrt{-\theta}}\right)$	$\frac{t^*}{2} \cdot \left( \left( \right) \right)$			
TEM	$\theta_1$			$\theta_{\rm r}$	
1D t-exponential		$2i\left[\frac{1}{(- heta_i)^{2-t^*}} ight]$		$-\mathbb{E}_i\left[(-\theta_i)^{2-t^*}\right]$	
1D <i>t</i> -Gaussian $(\mu = 0)$	$\left  -\mathbb{E}_{i} \left[ rac{1}{(- heta_{i})^{rac{1-t^{st}}{2}}}  ight] /\mathbb{E}$	$\left[\frac{1}{(- heta_i)^{rac{3-t^*}{2}}} ight]$	$-\frac{1}{(c_t*)}$	$rac{1}{(t^*)^{1-t^*}} \cdot \mathbb{E}_i\left[(- heta_i)^{rac{3-t^*}{2}} ight]$	

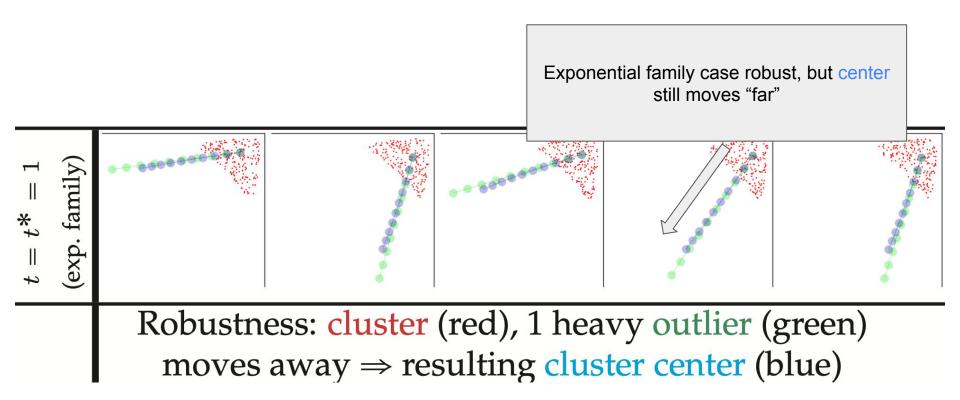
## Experiments

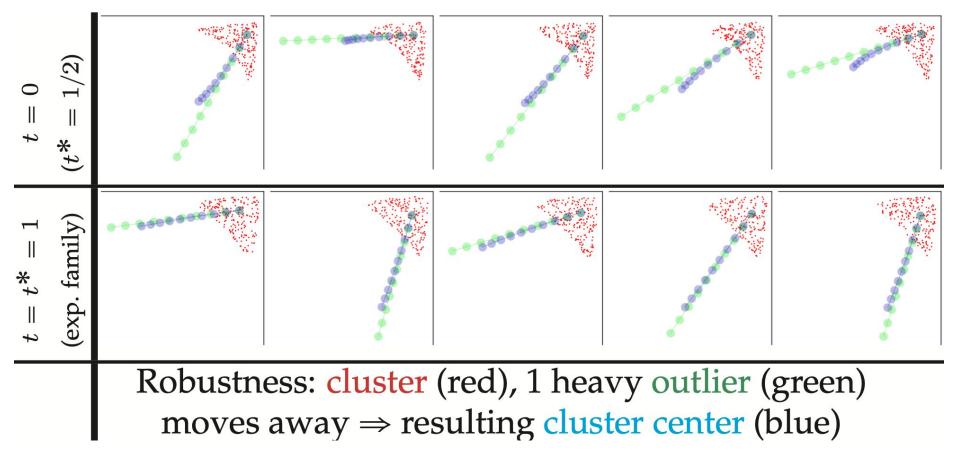
(more in paper)

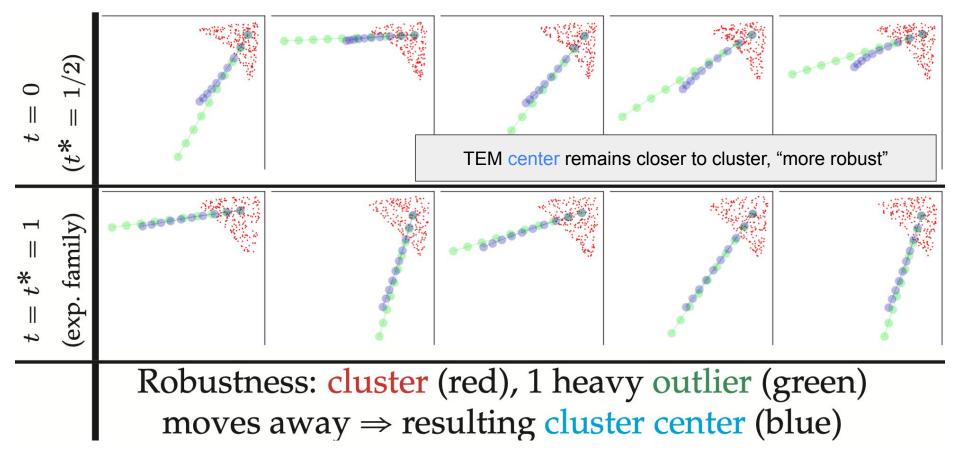
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# **Thank You**



Ehsan Amid



Richard Nock



Google Research

