

SMOOTHING OF NOISY  
RANDOM-TELEGRAPH-TYPE SIGNALS

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Abstract

For a class of Markov processes in additive gaussian white noise, both the equations for the optimal fixed-point smoother and those for a sub-optimal fixed-lag smoother are presented. Simulation results for the random telegraph wave are discussed.

1. INTRODUCTION

It has been shown in the literature<sup>[1]</sup> that fixed-lag smoothers for both the continuous and discrete-time linear-gaussian problems exhibit the following desirable properties: (1) the expected value of the mean square error is a non-increasing function of the lag and (2) for a value of lag of the order of the dominant time constant of the associated (zero-lag) filter, practically all of the possible improvement that smoothing offers over filtering is attained. Further, for stationary time-invariant processes<sup>[2]</sup>, the ratio of the fixed-lag error to the filtering error decreases as the signal-to-noise ratio increases. Hence, for those applications where a time delay between the estimate of the state of a process and the state itself is acceptable, it is worthwhile looking at the possibility of using fixed-lag smoothing rather than filtering.

Intuitively, it seems plausible that these properties should in fact hold for a larger class of processes than that above. In this paper, we derive and examine properties of a sub-optimal

fixed-lag smoothed estimate of the states of a class of Markov processes in the presence of additive gaussian white noise. Simulation of the fixed-lag smoother for the special case of the random telegraph wave suggests that the above properties do in fact hold for the class of processes under study.

2. A NONLINEAR FILTERING RESULT

Suppose  $\{x(t), t \geq 0\}$  is a Markov process with  $n$  distinct states  $a_1, \dots, a_n$  and stationary transition probabilities  $p_{ij}(h) \triangleq \Pr\{x(t+h) = a_j | x(t) = a_i\}$  satisfying

$$p_{ij}(h) = \begin{cases} 1 - v_i h + o(h) & i=j \\ v_{ij} h + o(h) & i \neq j \end{cases} \quad (1)$$

where  $h \geq 0$  and  $v_i, v_{ij}$  are nonnegative constants such that  $\sum_{j=1}^n v_{ij} = v_i$ . Let a measurement process  $\{z(t), t \geq 0\}$  be defined by

$$dz(t) = h(x(t))dt + \beta(t)dw(t) \quad (2)$$

Here  $dw(t)$  is an increment in a Wiener process independent of  $x$  with a derivative of covariance  $\delta(t-\tau)$ ,  $\beta(t)$  is a continuously differentiable function bounded away from zero and  $h$  is a real-

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real-valued function with values  $h_1, \dots, h_n$ ; not necessarily distinct. Denote the  $z$  process over the interval  $[0, t]$  by  $Z_t$  and denote the filtered probabilities  $\Pr\{x(t) = a_i | Z_t\}$  by  $P_i(t)$ . Then, following Wonham<sup>[3]</sup>, it can be shown that these probabilities satisfy

$$dp_i(t) = [-v_i P_i(t) + \sum_{\substack{j=1 \\ j \neq i}}^n v_j P_j(t)] dt + \beta(t)^{-2} P_i(t) [h_i - \bar{h}(t)] [dz(t) - \bar{h}(t) dt] \quad (3)$$

with  $\bar{h}(t) = \sum_{i=1}^n h_i P_i(t)$ , the conditional expectation of  $h(x(t))$ . The initial condition for (3) is just the initial probability distribution for  $x(0)$ .

### 3. FIXED-POINT SMOOTHING

For each fixed  $t$ , define the augmented process  $(X(t|\tau), \tau \geq t)$  by  $X(t|\tau) = [x(\tau) \ x(t)]^T$ . This process has  $N = n^2$  states  $A_I$ ,  $I=1, \dots, N$ ; or in component form  $(a_i, a_j)$  with  $I = (i, j)$  and  $i, j=1, \dots, n$ . From the properties of  $x$ , it also follows that this new process is Markov and has stationary transition probabilities  $P_{IJ}(h)$  satisfying

$$P_{IJ}(h) = \begin{cases} 1 - V_I h + o(h) & I=J \\ V_{IJ} h + o(h) & I \neq J \end{cases} \quad (4)$$

for  $I = (i_1, j_1)$  and  $J = (i_2, j_2)$ . The nonnegative constants  $V_I, V_{IJ}$  are defined by  $V_I = v_{i_1}, V_{IJ} = \delta_{j_1 j_2} v_{i_1 i_2}$  and satisfy  $\sum_{\substack{j=1 \\ j \neq i}}^n V_{IJ} = V_I$ .

With  $h(X(t|\tau)) = x(\tau)$ , the measurements are given by

$$dz(\tau) = x(\tau) d\tau + \beta(\tau) dw(\tau), \quad \tau \geq t \quad (5)$$

Then denote the filtered probabilities of  $X$ ,  $\Pr\{X(t|\tau) = A_I | Z_\tau\}$  by  $P_I(t|\tau)$  and the fixed-point smoothing probabilities of  $x$ ,  $\Pr\{x(t) = a_j | Z_t\}$  by  $p_j(t|\tau)$ . Since  $p_j(t|\tau) = \sum_{i=1}^n P_i(t|\tau)$  with  $I = (i, j)$  we have the filtered probabilities of  $X$  completely determining the fixed-point probabilities of  $x$ . From Section 2 the former are easily found as the solution of the equations

$$dP_I(t|\tau) = [-v_I P_I(t|\tau) + \sum_{\substack{j=1 \\ j \neq I}}^N v_{jI} P_j(t|\tau)] d\tau + \beta(t)^{-2} P_I(t|\tau) [a_i - \bar{x}(\tau)] [dz(\tau) - \bar{x}(\tau) d\tau] \quad (6)$$

with the initial condition  $P_I(t|t) = \delta_{i,j} P_i(t)$ ,  $I = (i, j)$  and with  $\bar{x}(\tau) = \sum_{i=1}^n a_i P_i(t|\tau)$ , the conditional expectation of  $x(\tau)$ .

The above approach to the continuous-time fixed-point smoothing problem has been used previously for both the linear<sup>[4]</sup> and nonlinear<sup>[5]</sup> cases.

### RANDOM TELEGRAPH WAVE EXAMPLE

For a process with states  $+1, -1$  and switching parameter  $v$  with  $v_{+1} = v$  and  $v_{-1} = v$  for  $i, j = 1, 2$ , the fixed-point equations (6) are, with independent variables  $q, Z$  and  $D$ ,

$$dq(\tau) = -2vq(\tau) + \beta^{-2} [1 - q(\tau)^2] dI(\tau), \quad q(0) = 0 \quad (7)$$

$$dZ(t|\tau) = \beta^{-2} [2D(t|\tau) - 1 - Z(t|\tau)q(\tau)] dI(\tau), \quad (8) \\ Z(t|t) = q(t)$$

$$dD(t|\tau) = v[1 - 2D(t|\tau)] d\tau + \beta^{-2} [\frac{1}{2}q(\tau) + \frac{1}{2}Z(t|\tau) - D(t|\tau)q(\tau)] dI(\tau), \\ D(t|t) = 1 \quad (9)$$

where  $q$  is the conditional expectation of  $x$ ,  $Z$  is the fixed-point estimate of  $x$ ,  $D$  is the probability that both components of the  $X$ -process are the same; and  $I$  is the innovations process with  $dI(\tau) = dz(\tau) - q(\tau)d\tau$ .

### 4. FIXED-LAG SMOOTHING (DISCRETE-TIME)

For the fixed-lag smoothing problem we need to know the evolution of the quantities  $P_I(t|t+L)$  with  $L$  fixed and  $t$  varying. One approach which suggests itself is to take equation (6) for  $P_I(t|\tau)$ , to express  $P_I(t|t+L)$  as an integral using this equation, and then to compute the differential now letting  $t$  rather than  $\tau$  vary. Although this method works for the linear-gaussian problem<sup>[6]</sup> it fails here because of the conditional rather than unconditional expectations in the integrand. We thus treat only the discrete-time case and derive a

suboptimal fixed-lag estimate as a linear combination of the states of a discrete-time nonlinear system driven by the measurements.

With obvious notation we have, discretizing (6)

$$\begin{aligned}
 P_I(k|\ell+1) &= P_I(k|\ell) + [-v_I P_I(k|\ell) + \sum_{\substack{J=1 \\ J \neq I}}^N v_{JI} P_J(k|\ell)]T \\
 &+ \beta(\ell)^{-2} P_I(k|\ell) [a_I - \bar{x}(\ell)] I(\ell) \quad (10)
 \end{aligned}$$

where  $k$  is fixed,  $\ell$  takes the values  $k, k+1, k+2, \dots$ , and  $I(\ell) = z(\ell+1) - z(\ell) - \bar{x}(\ell)T$ . Now, fix  $\ell$  in these equations, and in writing down (10) for each  $k$  in the range  $\ell$  to  $\ell-L+1$ , define the new variables  $F_{Ij}(\ell) = P_I(\ell-j+1|\ell)$  for  $I=1, \dots, N$  and  $j=1, \dots, L+1$ . We then obtain

$$\begin{aligned}
 F_{I,j+1}(\ell) &= F_{Ij}(\ell) + [-v_I F_{Ij}(\ell) + \sum_{\substack{J=1 \\ J \neq I}}^N v_{JI} F_{Jj}(\ell)]T \\
 &+ \beta(\ell)^{-2} F_{Ij}(\ell) [a_I - \bar{x}(\ell)] I(\ell) \quad (11) \\
 &\text{for } j=1, \dots, L
 \end{aligned}$$

while the quantities  $F_{11}(\ell)$  are none other than the filtered probabilities associated with the original  $x$ -process, and hence are updated by the discrete-time version of the filter equations (3). Finally, the suboptimal fixed-lag estimate of  $x(\ell-L)$  given  $Z_\ell$  is  $\sum_{I=1}^N a_I F_{I,L+1}(\ell)$ .

#### ALTERNATIVE DERIVATION

We now outline an alternative derivation of (11) which is based on a discrete-time approximation [7] to the continuous-time processes involved and the idea [8] that the discrete-time fixed-lag smoothing problem can be posed as one of filtering a related process. Consider first a discrete-time Markov process  $\{s(k), k=0, 1, \dots\}$  for which the transition probabilities  $\Pr\{s(k+1) = a_j | s(k) = a_i\}$  are denoted by  $P_{ij}$  with  $P_{ij} = 1 - v_i T$  for  $i=j$  and  $P_{ij} = v_{ij} T$  for  $i \neq j$ ,  $T$  being the discretization interval; and second, the observation process defined as

$$z(k) = s(k)T + \beta(k)w(k)\sqrt{T}$$

where  $\{w(k), k=0, 1, \dots\}$  is gaussian white noise

with unit variance.

Suppose we define a new Markov process  $\{S(k), k=L, L+1, \dots\}$  with  $S(k) = [s(k) \ s(k-1) \ \dots \ s(k-L)]^T$ , then the fixed-lag probabilities of the  $s$ -process are simply sums of the filtered probabilities of the  $S$ -process. The latter can be derived using a well-known recursive formula [9, p174] and approximating [7] for small  $T$ .

#### 5. SIMULATION - DISCUSSION AND RESULTS

For our example, the random telegraph wave, we generated the approximated discrete-time processes mentioned in the final part of section 4. It was found that for reliable results the following points need to be considered.

(a) Consider the discrete-time filter equation (see (7))

$$\begin{aligned}
 q(k+1) - q(k) &= -2vq(k)T + \beta^{-2}[1 - q(k)^2][x(k) - q(k)]T \\
 &+ \beta^{-1}[1 - q(k)^2]w(k)\sqrt{T}, \quad q(0) = 0 \quad (12)
 \end{aligned}$$

The definition of  $q$  implies  $|q(k)| \leq 1$  for all  $k$ . However, from (12) it is clear that for a sufficiently large value of  $w$ , this bound will be violated. Then, the term  $[1 - q(k)^2][x(k) - q(k)]$  is destabilizing in the sense that its effect on the increment is to cause  $|q(k+1)| > |q(k)|$ . Hence, for a sequence of values of  $w$  with suitable sign,  $q$  may become arbitrarily large. An ad hoc solution, which was found to be satisfactory, was to redefine  $q(k)$  as  $\pm 1$  at each iteration whenever  $|q(k)| > 1$ . An alternative method is to bound the noise samples and choose a sufficiently small sampling interval  $T$ .

(b) It is also necessary to ensure that the discrete-time processes are, in fact, approximations to their continuous-time counterparts. Each term on the right side of (12) is small if we choose  $T$  such that  $T < \beta/20\sqrt{v}$  and  $vT$  is small. Although ad hoc, this bounding proves to be a successful guideline in the choice of parameters in the simulation. This bound on  $T$  implies that the larger is the signal-to-noise ratio the greater is the amount of computation per unit of time.

An example of the results of simulation is shown in Fig. 1, with  $\nu = 50$ ,  $\beta = 0.07$  and  $T = 0.0005$ . Then  $\nu T = 0.025$  which is "small" and  $\beta/20\sqrt{\nu} \approx 0.0005$ . Thus the conditions on  $T$  are satisfied. It is apparent from Fig. 1 that there is an initial rapid improvement in estimation with lag, followed by an evening out to a steady-state value. For the filter eq.(12), we can associate a time constant of  $1/2\nu T$  sampling intervals. For our example,  $1/2\nu T = 20$ . Compare this with Fig. 1 where the lag for which practically all improvement is obtained is about 15. We found that every example for which the conditions on  $\beta$ ,  $\nu$  and  $T$  were satisfied, this type of behaviour was observed.

Fig. 2 shows a plot of improvement of smoothing over filtering versus a noise measure  $\mu = \nu\beta^2$ . The similarity with the linear-gaussian case can be seen.

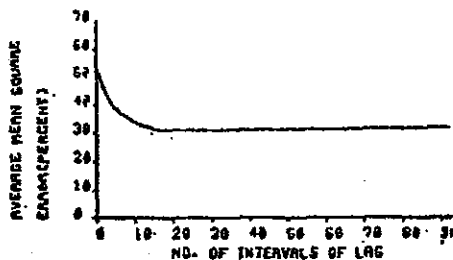


Figure 1

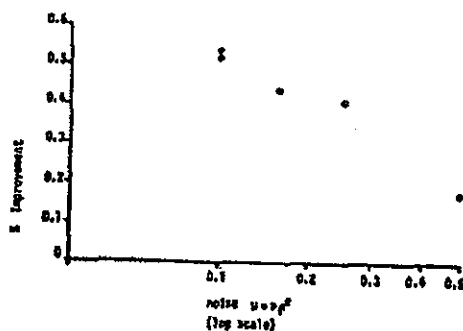


Figure 2

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